Week 3

Introduction to Likelihood Inference

University of New Orleans
POLI 6003
Rich Frank
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Altman and Brehm (2003) takeaways

- Details are important.
- Software default settings can have an effect.
 - Always keep track of software and version used.
- Rounding/truncating
- Random number generators
 - Will become important when we discuss out-ofsample predictions of marginal effects.

- Even pros make mistakes.
- Ability to replicate is essential.

Distributions

• We have some uncertainty as to how the data were generated.

• We have to be clear as to what extent we think the sample represents the population as well as our uncertainty about the data generation process.

Linking DVs, distributions, and models

- King (1989) spends a chapter introducing a number of distributions that we will be seeing this semester.
 - Bernoulli
 - Binomial
 - Extended beta-binomial
 - Poisson
 - Negative binomial
 - Normal
 - Log-normal

Probability

Probability and likelihood differ from one another principally by how they treat the data and model in relation to one another.

 Probability theory presumes some given model (or set of parameters) and seeks to estimate the data (given those parameters).

As King (1989:9) puts it:

$$\mathbf{Y} \sim f(\mathbf{y}|\boldsymbol{\theta}, \boldsymbol{\alpha})$$

And

$$\theta = g(X, \beta)$$

• So our data, Y, has a probability distribution given by parameters θ and α , and θ is a function of some variables X and their parameters, β .

 All this comprise King's model so the normal probability statement appears is:

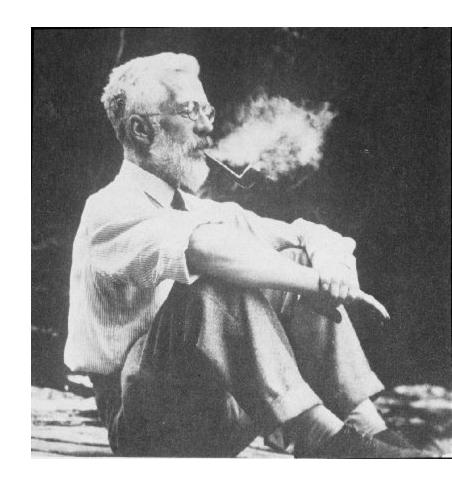
$$Pr(y|M) \equiv Pr(\text{data}|\text{model})$$

- Some problems with using probability model:
 - 1. Presumes that data are random and unknown.
 - 2. Assumes that the model is known.

• What would be useful to have is the inverse probability: $Pr(M \mid y)$

- But that requires knowledge or strong assumptions about the important elements of the unknown model, θ .
- Bayesian methods make some (often weak) assumptions about the model given what we know about the world, but these are beyond the scope of this class.
- To a certain extent complexity theory also makes similar assumptions to generate data.

- As a result likelihood methods of inference have become popular.
- Attributed to the mathematician R.A.
 Fischer.
- Likelihood methods estimate the model given the data.



Likelihood depends on the following axiom:

$$L(\tilde{\theta} \mid y, M^*) \equiv L(\tilde{\theta} \mid y)$$
$$= k(y)Pr(y|\tilde{\theta})$$
$$\alpha Pr(y|\tilde{\theta})$$

- Where $\tilde{\theta}$ represents the hypothetical value of θ
- k(y) = the constant of proportionality, which is constant across all hypothetical $\tilde{\theta}$ but represents the way (the functional form) that the data shape $\tilde{\theta}$. This enables us to estimate likelihood as a measure of relative (rather than absolute) uncertainty.

• Thus likelihood is proportional to traditional probability where the constant k(y) is an unknown function of the data.

• The uncertainty is relative to other possible functions of y and the hypothetical values of $\tilde{\theta}$.

• Therefore, it measures the relative likelihood of a specific hypothetical $\tilde{\beta}$ producing the data we observe.

Examples

Let's think about this a bit less theoretically.

 Okay, so the goal of likelihood is to estimate parameters given the data.

- Consider the data fixed.
- And consider a distribution of parameters $\tilde{\theta}$ we want to find the $\tilde{\theta}$ that is most likely to have generated the data we see.

Suppose we have a model where

$$Y \sim N(\mu, \sigma^2)$$

 $E(Y) = \mu$
 $Var(Y) = \sigma^2$

- So the model presumes the data are normally distributed.
- We want to estimate the parameters μ and σ^2
- We want to find values of the mean and variance that are most likely to have produced the data, *Y*.

■ Imagine that we have data on annual measures of presidential approval over an 8 year period.

$$Y = [54 53 49 61 58 62 50 52]'$$

- We want to know the chances that the data are drawn from a distribution of mean μ and variance σ^2 .
- What do you think is the likely mean of the distribution?
- Maximum likelihood is a more formal and systematic way of finding the parameters of the distribution most likely to have generated the data.

• If we assume that the data are normally distributed, then the PDF is given by:

$$Pr(Y=y_i) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(y_i - \mu_i)^2}{2\sigma^2}\right]}$$

So we can compute the probability of any particular observation in the distribution by solving the equation using that value from the data.

- However we are more interested in the *joint* probability of all of the 8 observations of presidential approval rather than just 1 from a distribution with a particular mean and variance.
- Assuming that the observations are independent of each other (a leap in this case) the joint PDF is equal to the product of the marginal probabilities.

$$Pr(A \text{ and } B) = Pr(A) \cdot Pr(B)$$

So the joint probability is given by:

$$Pr(Y) = y_i \forall i) = L(Y|\mu, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(y_i - \mu_i)^2}{2\sigma^2}\right]}$$

- This formula assumes the parameters are given while we want to estimate them.
- Fortunately the likelihood of the parameters is proportional to the probability of the data given the parameters.

$$L(\mu, \sigma^{2} | \mathbf{Y}) = k(y) \prod_{i=1}^{n} f_{normal}(\mathbf{Y} | \mu, \sigma^{2})$$

$$\propto \prod_{i=1}^{n} f_{normal}(\mathbf{Y} | \mu, \sigma^{2})$$

$$= \prod_{i=1}^{n} \frac{1}{\sqrt{2\pi\sigma^{2}}} e^{\left[\frac{-(y_{i} - \mu_{i})^{2}}{2\sigma^{2}}\right]}$$

- This is the likelihood function, the constant k(y) ensures that the probability and the likelihood are proportional (\propto).
- Now we can use this function to establish which values of the mean and variance are most likely to have generated the data.

- We can do this by hand (at least just this once).
- First, let's pick a value of μ , and for convenience set σ^2 to 1. (King tells us that this is the stylized Normal distribution—it relies on the independence of the mean and variance in the Normal distribution).

• Let $\mu = 53$.

$$L = \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(54-53)^2}{2\sigma^2}\right]} \cdot \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(53-53)^2}{2\sigma^2}\right]}$$

$$\frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(49-53)^2}{2\sigma^2}\right]} \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(61-53)^2}{2\sigma^2}\right]}$$

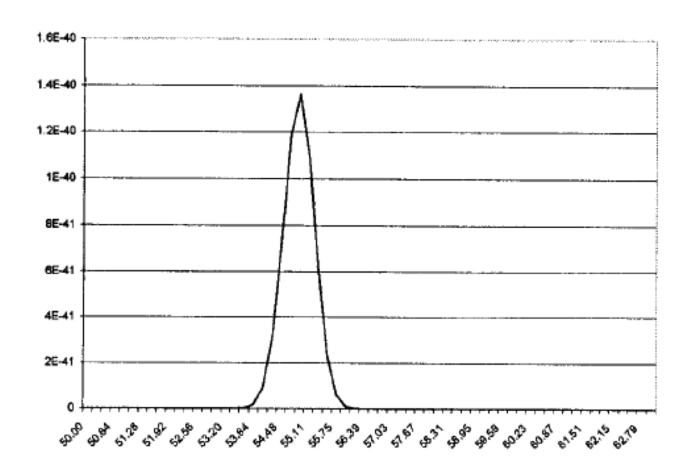
$$\frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(58-53)^2}{2\sigma^2}\right]} \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(62-53)^2}{2\sigma^2}\right]}$$

$$\frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(50-53)^2}{2\sigma^2}\right]} \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(52-53)^2}{2\sigma^2}\right]} \approx 1.752e-45$$

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- That is a really small number. Thus we have little reason to think that 53 is the mean of the distribution that generated the data.
- So what we want to do is to do this same calculation for a number of different possible values of μ .
- The largest of the likelihoods will be the maximum of the likelihood function.
- The most efficient way to represent this is in a figure.



Likelihood estimate of mean presidential approval

- From the figure it looks like the maximum is approximately 55. This is the ML estimate of μ .
- Another way of saying it is that the value of μ that maximizes the likelihood function is 55.

- As you can see this is a bit clunky to do, partly because products are harder to deal with than sums.
- Fortunately, we can transform the likelihood function above by any monotonic form (like the natural log).

$$\ln L(\mu, \sigma^2 | \mathbf{Y}) = \ln \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(y_i - \mu_i)^2}{2\sigma^2}\right]}$$

$$= \sum \ln \left[\frac{1}{\sqrt{2\pi\sigma^2}} e^{\left[\frac{-(y_i - \mu_i)^2}{2\sigma^2} \right]} \right]$$

$$= -\frac{1}{2}(\ln(2\pi)) - \frac{1}{2}(\ln(\sigma^2)) - \frac{1}{2\sigma^2} \left[\sum_{i=1}^n (y_i - u_i)^2\right]$$

 Using this log-likelihood function we conduct similar calculations to what we did above, plot the estimates, and visually find the maximum.

 This method is similar to what numerical methods do without the graphics.

- Let's try a different example.
- Suppose we have data on 20 states about whether they have adopted a lottery.

$$Y = [0000000000011111111111]'$$

■ The probit likelihood function is given by:

$$ln L = \sum_{i=1}^{n} y_i \ln(\Phi(X\beta)) + (1 - y_i) \ln(1 - \Phi(X\beta))$$

- This is doable. For each of the 20 observations (y_i) we multiply y_i by the log of $\Phi(X\beta)$ and then add 1- y_i multiplied by the log of 1- $\Phi(X\beta)$.
- Φ(Xβ) would be the independent variables multiplied by their coefficients, summed (a z-score) and then evaluated on the normal CDF thus giving us a probability.
- However, we have no independent variables in this example.
- Instead, like the above example we will test different values of probability as the mean probability responsible for generating the data on lottery adoption.

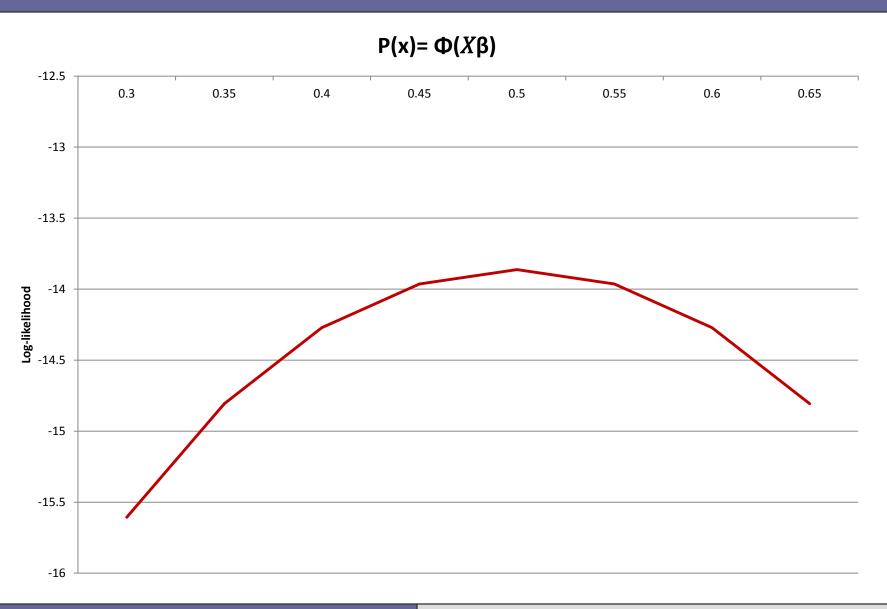
	p(z)=.3	p(z)=.35	p(z)=.4	p(z)=.45	p(z)=.5	p(z)=.55	p(z)=.6	p(z)=.65
Υ	0.3	.35	.4	.45	.5	.55	.6	.65
0	356675	-0.430783	510826	597837	693147	798508	916291	-1.049822
0	356675	-0.430783	510826	597837	693147	798508	916291	-1.049822
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1	-1.203973	-1.049822	916291	798508	693147	597837	510826	430783
1	-1.203973	-1.049822	916291	798508	693147	597837	510826	430783
1	-1.203973	-1.049822	916291	798508	693147	597837	510826	430783
1	-1.203973	-1.049822	916291	798508	693147	597837	510826	430783
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1	-1.203973	-1.049822	916291	798508	693147	597837	510826	430783
1	-1.203973	-1.049822	916291	798508	693147	597837	510826	430783
SUM	-15.60648	-14.80605	-14.27116	-13.96345	-13.86294	-13.96345	-14.27116	-14.80605

For example for the first column

$$\ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 0 * \ln(.3) + (1-0) * \ln(1-.3) + \\ \ln L = 1 * \ln(.3) + (1-1) * \ln(1-$$

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Likelihood estimates of lottery adoption



MLE

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■ So the maximum is at .5.

■ This should not be a surprise because out of the 20 observations there are 10 zeros and 10 ones.

Plug and chug iterative search

- So now we have a rough intuition as to what your software is doing when it is estimating an ML model (e.g. logit).
- The software takes a starting value of β (either zero or an OLS estimate) to estimate the log-likelihood (LL).
- It takes the first derivative of the LL with respect of the parameters to find the gradient.
- The gradient tells us the slope of a line tangent to the curve at the point of the LL estimate.
- If the gradient is positive then the L is increasing in β .
- It then increases the estimate of β and try again.
 - If the slope is negative it decreases the estimate.

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 Once the first derivative is approaching zero, it stops and looks at the second derivative.

• If it is negative then it has reached a maximum.

• The flatter the slope the harder it can be to determine that we are at the top.

■ The second derivative tells us how quickly the slope is changing, so we know how large of a step to take.

Properties of MLEs

- Consistency—They are asymptotically consistent. As sample size increases, the estimates increasingly resemble the actual population parameters. As a result MLEs are good large sample estimators (what is large?)
- Asymptotic normalcy—The MLE parameters are distributed according to the standard multivariate normal no mater what distribution assumptions you make in your model. This allows us to describe them using z-scores.

Properties of MLEs

• Asymptotic efficiency—basically this means that MLE has the smallest asymptotic variance of any estimators that are also consistent and asymptotically normal.

■ **Invariance**—If Θ_{ML} is a vector of ML estimates, and $g(\Theta)$ is a continuous function of Θ , then $g(\Theta_{ML})$ is a consistent estimator of $g(\Theta)$. So if we transform variables, we can retransform the estimates without losing interpretive ability.

• Again, we can only speak about <u>relative</u> likelihood not <u>absolute</u> likelihood of our estimates given a set of data.

 Next week, we will begin to talk about likelihood inference applied to binary dependent variables.